Introduction to imprecise probabilities SIPTA School 2024, Ghent, Belgium

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12 August 2024

Overview

Kick-off (slot 1)

Classical probability theory (slot 1)

Interpretation of probability (slot 2)

Limitations of probability theory (slot 2)

Probability intervals (slot 3)

Credal sets (slot 3-4)

Acceptability & Desirability (slot 4–5)

Interval expectation & probability (slot 5–6)

Overview

Kick-off (slot 1) Uncertainty, representation & reasoning Illustration: Blood groups

Classical probability theory (slot 1)

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Probability intervals (slot 3)
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Acceptability & Desirability (slot 4–5)

Interval expectation & probability (slot 5-6)

What is *uncertainty*?

There is no consensus definition of uncertainty...



Uncertainty is the lack of certainty, a state of limited knowledge where it is impossible to exactly describe the existing state, a future outcome, or more than one possible outcome.

Hubbard (2014)

How can we represent uncertainty?

Uncertainty representations are defined by axioms

Any reasonable measure of belief is isomorphic to a probability distribution.

Cox (1946), paraphrased

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Some concrete examples of representations

- probability distributions
- ▶ intervals, sets
- possibility distributions, belief functions
- sets of desirable gambles, preference orders
- Iower/upper probabilities/expectations
- credal sets
- choice functions

What is *reasoning* under uncertainty?

Reasoning under uncertainty: deducing inferences or making decisions

In each problem situation, uncertainty is associated with some purpose. It may, for example, be associated with prediction, retrodiction, prescription, or decision making.



Reasoning under uncertainty: deducing inferences or making decisions

In each problem situation, uncertainty is associated with some purpose. It may, for example, be associated with prediction, retrodiction, prescription, or decision making.

Some concrete examples

- Predicting individual disease risk based on medical history
- Planning and designing of blood bank for a hospital with 100 beds
- A newspaper vendor must decide how many copies to purchase each day

Klir (1995)

Problem setup

- ► A sample:
 - A AB O B A B A O
- A disease
- Two treatments with differing effectiveness:

	А	В	AB	0
f_1	0.5	0.6	0.7	0.1
f_2	0.4	0.3	0.3	0.8

Problem setup

Representation

- ► A sample:
 - A AB O B A B A O
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- Two treatments with differing effectiveness:

 $\begin{array}{ccccccc} A & B & AB & O \\ f_1 & 0.5 & 0.6 & 0.7 & 0.1 \\ f_2 & 0.4 & 0.3 & 0.3 & 0.8 \end{array}$

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Representation

PMF from observed frequencies:

p_{A}	p_{B}	p_{AB}	$p_{\rm O}$
3/8	2/8	1/8	2/8

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Reasoning

- Probability P({A, O})?
- Expectation $E_p(f_1)$?
- Outcome with maximal probability?
- Treatment with highest expected effectiveness?

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Representation

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Reasoning

- $P(\{A, O\}) = 5/8$
- $E_p(f_1) = 0.45 \ (check!)$
- Outcome A has maximal probability
- Treatment 2 has highest expected effectiveness (check!)

Problem setup

- ► A partial information sample:
 - A,B AB O A,B A,B A,B A,B O
- A disease
- Two treatments with differing effectiveness:

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Representation

Blood groups problem with cheap test

Problem setup

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Representation

PMF from observed frequencies:

p_{A}	p_{B}	p_{AB}	$p_{\rm O}$
?	?	1/8	2/8

Reasoning

- ► $P({A, O}) = ?$
- $\blacktriangleright E_p(f_1) = ?$
- Outcome ? has maximal probability
- Treatment ? has highest expected effectiveness

We want to be able to represent uncertainty and reason also in situations with partial information

Overview

Kick-off (slot 1)

Classical probability theory (slot 1) Representation Reasoning Learning

Multivariate probability theory

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Basic setup and axioms

Basic setup of probability theory:

- Random variable X
- ▶ Set of possible outcomes X
- ► Each possible event/set S ⊆ X is assigned a probability value P(S)

Basic setup and axioms

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Basic setup and axioms

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Axioms

A probability mass function p (and the corresponding probability measure P) must be:

- 1. Nonnegative: $p_x \ge 0$ for all outcomes $x \in \mathcal{X}$
- 2. Additive: $P(S) = \sum_{x \in S} p_x$ for all events $S \subseteq \mathcal{X}$
- 3. Normed: $P(\mathcal{X}) = 1$



Probability simplex: visualizing probability mass functions



Probability simplex: visualizing probability mass functions



Probability simplex: visualizing probability mass functions

- $\blacktriangleright \mathcal{X} = \{\mathsf{Win}, \mathsf{Draw}, \mathsf{Lose}\}$
- 'degenerate' probability mass functions (pmfs)
 p = (p_W, p_D, p_L) at the corners
- other pmfs as convex combinations thereof; values can be 'read off' as distance to opposite edge



Probability simplex: visualizing probability mass functions — exercise

The pmf

► The pmf

Win Lose On a gridded probability simplex on vour worksheet, indicate The degenerate pmf corresponding to $p_{\rm l} = 1$ $(p_{\rm W}, p_{\rm D}, p_{\rm L}) = (05, 0.5, 0)$ $(p_{\rm W}, p_{\rm D}, p_{\rm L}) = (0.1, 0.3, 0.6)$ Draw

Reasoning — deducing inferences and making decisions

Deducing inferences

probability values

• expectations/previsions of real-valued functions f on \mathcal{X} :

$$E_p(f) := \sum_{x \in \mathcal{X}} p_x f(x)$$

Reasoning — deducing inferences and making decisions

Deducing inferences

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Decision making

- outcomes with maximal probability
- options minimizing/maximizing expectation
- Topic of tomorrow morning's lecture

Probability simplex: visualizing probabilities and expectations

X = {Win, Draw, Lose}
Visualize probability and expectation values as the lines of pmfs for which that value is attained



Lose

Probability simplex: visualizing probabilities and expectations

- $\blacktriangleright \ \mathcal{X} = \{\mathsf{Win}, \mathsf{Draw}, \mathsf{Lose}\}$
- Visualize probability and expectation values as the lines of pmfs for which that value is attained

Probability example

- $A = \{Win, Lose\}$
- ► *P*(*A*) = 2/3



Probability simplex: visualizing probabilities and expectations

- $\blacktriangleright \mathcal{X} = \{\mathsf{Win}, \mathsf{Draw}, \mathsf{Lose}\}$
- Visualize probability and expectation values as the lines of pmfs for which that value is attained

Probability example

- $A = \{Win, Lose\}$
- ▶ $P(A) = p_W + p_L = 2/3$


Probability simplex: visualizing probabilities and expectations

- $\blacktriangleright \mathcal{X} = \{\mathsf{Win}, \mathsf{Draw}, \mathsf{Lose}\}$
- Visualize probability and expectation values as the lines of pmfs for which that value is attained

Expectation example

- ► $f = (f_W, f_D, f_L) = (1, 0, -1)$
- E(f) = -1/2



Probability simplex: visualizing probabilities and expectations

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- Visualize probability and expectation values as the lines of pmfs for which that value is attained

Expectation example

- $f = (f_W, f_D, f_L) = (1, 0, -1)$
- $E(f) = p_W p_L = -1/2$



Probability simplex: visualizing probabilities and expectations — exercise

Win

On a non-gridded probability simplex on your worksheet, indicate

- The set of pmfs for which P({D,L}) = 0.4
- The set of pmfs for which E(f) = 0 with f = (-1, -4, 4)
- Is there a pmf compatible with both?

(If yes, which?)



Learning — creating a representation

from data, using estimation techniques (learning)

from experts, using elicitation (asking questions)

Multivariate probability mass functions: basic setup

• Index set
$$N = \{1, \ldots, n\}$$

• Multivariate variable
$$\boldsymbol{X} = (X_1, X_2, \dots, X_n)$$

- ▶ Set of possible outcomes $\mathbf{x} \in \mathcal{X} = \mathcal{X}_1 \times \mathcal{X}_2 \times \ldots \times \mathcal{X}_n$
- Each possible outcome is assigned a probability value

Marginal probability

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Marginal probabilities are probabilities defined for events corresponding to setting some components of the variable to given values

• Let
$$K \subseteq N$$
 and
 $m{X}_K := (X_k : k \in K)$
 $m{x}_K \in \mathcal{X}_K := \bigotimes_{k \in K} \mathcal{X}_k$

Marginal probabilities follow from the additivity axiom:

$$P(\boldsymbol{X}_{K} = \boldsymbol{x}_{K}) = \sum_{\substack{\boldsymbol{z} \in \mathcal{X} \\ \boldsymbol{z}_{K} = \boldsymbol{x}_{K}}} p_{\boldsymbol{z}}$$

Marginal probability example

Problem setup



Marginal probability example

Problem setup

$\blacktriangleright \mathbf{X} = (X_1, X_2, X_3)$													
$\mathcal{X}_1 =$	= {A,	B, C},	, $X_2 =$	= {0, 1	.}, X	$i_3 = \{$	+, -	}					
x	x_1 x_2 x_3	A 0 +	A 0 —	A 1 +	A 1 -	B 0 +	B 0	B 1 +	B 1 —	C 0 +	C 0 —	C 1 +	C 1 —
p_{x}		0.2	0.1	0.1	0	0.1	0	0	0.3	0	0	0	0.2

Inferences

► $P(X_1 = B)?$

• $P(X_1 = A, X_2 = 0)?$

►
$$P(X_1 = C, X_3 = +)?$$

Marginal probability example

Problem setup

	$\blacktriangleright \mathbf{X} = (X_1, X_2, X_3)$													
	$\mathcal{X}_1 =$	= {A,	B, C},	, $X_2 =$	= {0, 1	.}, X	$i_3 = \{$	+, -	}					
•	x	x_1 x_2 x_3	A 0 +	A 0 —	A 1 +	A 1 —	B 0 +	B 0	B 1 +	B 1 —	C 0 +	C 0	C 1 +	C 1
	p_{x}		0.2	0.1	0.1	0	0.1	0	0	0.3	0	0	0	0.2

Inferences

- $P(X_1 = B) = 0.4$
- $P(X_1 = A, X_2 = 0) = 0.3$

►
$$P(X_1 = C, X_3 = +) = 0$$

Conditional probabilities

- Conditional probabilities are probabilities that hold
 - assuming some event is known to be true, or specifically
 - assuming some random variables take some given values

▶ Let
$$B \subseteq \mathcal{X}$$
 for which $P(B) > 0$, then for $A \subseteq \mathcal{X}$ we have $P(A|B) = \frac{P(A \cap B)}{P(B)}$

• Let $K \subseteq N$ and

$$oldsymbol{X}_{K} := (X_k : k \in K)$$

 $oldsymbol{x}_{K} \in \mathcal{X}_{K} := oldsymbol{X}_{K} \in \mathcal{X}_{k}$

with $P(\boldsymbol{X}_{K} = \boldsymbol{x}_{K}) > 0$ then

$$P(\boldsymbol{X}_{N\setminus K} = \boldsymbol{x}_{N\setminus K} | \boldsymbol{X}_{K} = \boldsymbol{x}_{K}) = \frac{P(\boldsymbol{X}_{K} = \boldsymbol{x}_{K}, \boldsymbol{X}_{N\setminus K} = \boldsymbol{x}_{N\setminus K})}{P(\boldsymbol{X}_{K} = \boldsymbol{x}_{K})}$$

Conditional probability example

Problem setup



Conditional probability example

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$\blacktriangleright \mathbf{X} = (X_1, X_2, X_3)$													
$\mathcal{X}_1 =$: {A,	B, C}	, X ₂ =	= {0, 1	.}, X	$i_3 = \{$	+, -	}					
x	x ₁ x ₂ x ₃	A 0 +	A 0 —	A 1 +	A 1 -	B 0 +	B 0	B 1 +	B 1 —	C 0 +	C 0 —	C 1 +	C 1 —
p_x		0.2	0.1	0.1	0	0.1	0	0	0.3	0	0	0	0.2

Inferences

- $P(X_1 = A | X_2 = 0, X_3 = +)?$
- $P(X_1 = B | X_2 = 0, X_3 = -)?$
- $P(X_2 = 1 | X_1 = C, X_3 = +)?$

Conditional probability example

Problem setup

$\blacktriangleright \mathbf{X} = (X_1, X_2, X_3)$													
$\mathcal{X}_1 =$: {A,	B, C}	, X ₂ =	= {0, 1	.}, X	$i_3 = \{$	+, -	}					
x	x ₁ x ₂ x ₃	A 0 +	A 0 —	A 1 +	A 1 -	B 0 +	B 0	B 1 +	B 1 —	C 0 +	C 0 —	C 1 +	C 1 —
p_x		0.2	0.1	0.1	0	0.1	0	0	0.3	0	0	0	0.2

Inferences

►
$$P(X_1 = A | X_2 = 0, X_3 = +) = \frac{2}{3}$$

►
$$P(X_1 = B | X_2 = 0, X_3 = -) = 0$$

•
$$P(X_2 = 1 | X_1 = C, X_3 = +)$$
 is not well-defined (why?)

Independence

► Two random variables X₁ and X₂ are independent if their well-defined conditionals coincide with the marginals for all x ∈ X:

$$P(X_1 = x_1 | X_2 = x_2) = P(X_1 = x_1)$$

$$P(X_2 = x_2 | X_1 = x_1) = P(X_2 = x_2)$$

This is equivalent to the joint factorizing:

$$P(X_1 = x_1, X_2 = x_2) = P(X_1 = x_1)P(X_2 = x_2)$$
 for all $\boldsymbol{x} \in \mathcal{X}$

Reasoning — deducing inferences and making decisions

What can be done with (joint) probabilities can also be done with marginal and conditional probabilities

Reasoning — deducing inferences and making decisions

What can be done with (joint) probabilities can also be done with marginal and conditional probabilities

Learning — creating a representation

Marginal and conditional probabilities can be either

- deduced from the learned (joint) probabilities, or
- learned directly, to together define the joint, possibly using independence assumptions

Multivariate probability — exercise

Consider the (joint) random variable $\mathbf{X} = (X_1, X_2, X_3)$ with $\mathcal{X} = \{0, 1\}^3$. The joint probabilities are determined by the following information:

- \triangleright X₁ and X₂ are independent
- It holds that $P(X_1=0)=1/2$ and $P(X_2=0)=1/5$
- ▶ The conditional probabilities $P(X_3 | X_1, X_2)$ are determined by the following table:

$$\begin{array}{c|ccccc} x_1 & 0 & 0 & 1 & 1 \\ x_2 & 0 & 1 & 0 & 1 \\ \hline P(X_3 = 0 \,|\, X_1 = x_1, X_2 = x_2) & \frac{1}{3} & \frac{1}{4} & \frac{1}{5} & \frac{1}{6} \end{array}$$

Derive that $P(X_1 = 0 | X_2 = 0, X_3 = 0) = \frac{5}{8}$.

As part of your calculation, you should derive and write a general expression for $P(X_1 | X_2, X_3)$ in terms of the (symbolic) probabilities that are given—before filling in the specific numbers.

Overview

Kick-off (slot 1)

Classical probability theory (slot 1)

Interpretation of probability (slot 2) Overview of interpretations Betting game

Limitations of probability theory (slot 2)

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Credal sets (slot 3–4)

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A forecast states:

"There is an 80% probability of showers for tomorrow"

What does this mean?

(More generally, what is the meaning of probability values?)

Diversity in interpretations of probability

There are several schools of thought regarding the interpretation of probabilities, none of them without flaws, internal contradictions, or paradoxes.

De Elía & Laprise (2005)

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Diversity in interpretations of probability

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Major interpretations physical evidential graded belief

frequentist classical subjective propensity logical

Physical

- Concerns statements about events related to physical systems
- Connected to frequency of occurrence of these events

Evidential

- A measure for the evidence supporting some (any) statement
- Typically intended to be objective

Graded belief

- A degree of belief about some (any) statement
- Typically subjective

Kaplan (2014)

Frequentist (physical interpretation)

An event's probability is defined as the **limit** of its **relative frequency** in many trials (the long-run probability).

Kaplan (2014)

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How short or long can or should the sequence of trials be?

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Classical (evidential interpretation)

Assuming **equally possible** cases, the probability of an event is the **ratio** of the relative number of cases favorable to it.

Wikipedia (paraphrased)

Kaplan (2014)

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Kaplan (2014)

What does the assumption mean and what are its grounds?

Subjective (graded belief interpretation)

A subjective probability is anyone's **opinion** of what the probability is for an event.

Siegel & Wagner (2022)

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Is it a problem if probability values are opinions?

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Siegel & Wagner (2022)

Is it a problem if probability values are opinions?

Betting interpretation (special case of subjective interpretation)

- Probabilities are defined by the subject's betting behavior (therefore also called *behavioral interpretation*)
- Relevant for our discussion of imprecise probability theories

Betting game setup

- ► Two players:
 - subject (gambler)
 - **bookie** (proposes bets)
- **Gambles** from the subject's perspective:

$$1_x(z) = egin{cases} 1 & ext{if } z = x \ 0 & ext{if } z
eq x \end{cases}$$

for all $x \in \mathcal{X}$

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Eliciting probabilities

- For each of the gambles 1_x, the subject offers their fair price p_x
- ► The bookie can propose to exchange one for the other, i.e., 1_x - p_x or p_x - 1_x, which the subject is committed to accept
- The p_x are the subject's probabilities

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- ► The bookie can propose to exchange one for the other, i.e., 1_x - p_x or p_x - 1_x, which the subject is committed to accept
- The p_x are the subject's probabilities
- Should the subject state any set of fair prices?

Coherence: deriving the axioms of probability from the betting game

- Assume the gambler specifies a negative price $p_x < 0$.
- Exchange proposed by the bookie:

$$p_x - 1_x = egin{cases} p_x - 1 < 0 & ext{if } x ext{ occurs} \ p_x < 0 & ext{ otherwise} \end{cases}$$

- Should be unacceptable to the subject, because it implies a sure loss, i.e., is negative whatever occurs!
- Nonnegativity axiom is required

Coherence: deriving the axioms of probability from the betting game

The bookie can choose to propose multiple bets, concerning some $S \subseteq \mathcal{X}$.

Combined exchange proposed by the bookie:

$$\sum_{x \in S} (p_x - 1_x) = \left(\sum_{x \in S} p_x\right) - 1_S, \text{ with } 1_S = \sum_{x \in S} 1_x = \begin{cases} 1 & \text{if } S \text{ occurs} \\ 0 & \text{otherwise} \end{cases}$$

▶ The fair price P(S) for the combined bet is $\sum_{x \in S} p_x$

Additivity axiom follows

Coherence: deriving the axioms of probability from the betting game

- ▶ The bookie can choose to combine the bets for all $x \in \mathcal{X}$, so for the gamble $1_{\mathcal{X}} = 1$
- ▶ Possible exchanges to be proposed: P(X) 1 or 1 P(X)
- Exchanges do not depend on the $x \in \mathcal{X}$ that occurs
- ▶ If $P(X) \neq 1$, one possible proposal would result in a *sure loss*
- The normedness axiom is required

Overview

Kick-off (slot 1)

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Interpretation of probability (slot 2)

Limitations of probability theory (slot 2) Aleatoric vs. epistemic uncertainty Representing epistemic uncertainty Rational agents vs. Real agents Why go beyond probability theory? Probability intervals (slot 3)

Credal sets (slot 3–4)

Acceptability & Desirability (slot 4–5)

Interval expectation & probability (slot 5-6)
Aleatoric uncertainty — the irreducible part

- Aleatoric uncertainty arises from random variation
- Additional information cannot reduce it
- > Other names: variability, stochasticity, randomness, chance, risk

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Epistemic uncertainty — the reducible part

- Epistemic uncertainty arises from a lack of knowledge (also at inference time)
- Additional information can reduce or eliminate it
- > Other names: *incertitude*, *ambiguity*, *ignorance*, *imprecision*

Aleatoric uncertainty — the irreducible part

- Aleatoric uncertainty arises from random variation
- Additional information cannot reduce it
- > Other names: variability, stochasticity, randomness, chance, risk
- Sources: spatial variation, temporal fluctuations, manufacturing or genetic differences,...

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- Additional information can reduce or eliminate it
- > Other names: *incertitude*, *ambiguity*, *ignorance*, *imprecision*
- Sources: limited sample size, mensurational limits ('measurement error'), censoring, poorly defined outcomes,...

Outcome of toss of fair coin

Epistemic uncertainty examples

Bias of a coin for tossing

- Outcome of toss of fair coin
- Decay time of a radioactive atom

Epistemic uncertainty examples

- Bias of a coin for tossing
- Weight of a proton

- Outcome of toss of fair coin
- Decay time of a radioactive atom
- Value of a decimal of a randomly generated number

Epistemic uncertainty examples

- Bias of a coin for tossing
- Weight of a proton
- Value of a decimal of an irrational mathematical constant

- Outcome of toss of fair coin
- Decay time of a radioactive atom
- Value of a decimal of a randomly generated number
- 'Noise'

Epistemic uncertainty examples

- Bias of a coin for tossing
- Weight of a proton
- Value of a decimal of an irrational mathematical constant
- Model uncertainty
 - Parameter values
 - Dependencies
 - Functional forms
 - Level of abstraction

Can probability theory differentiate between aleatoric and epistemic uncertainty?

Two different coins

Coin	S	L
Flips	2	$2\cdot 10^6$
Heads	50%	50%
Tails	50%	50%

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- What can you say about the reliability of the estimate for each coin?
- How would you communicate the probability estimates?
- How would you communicate derived inferences and decisions?

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for both coins

- What can you say about the reliability of the estimate for each coin?
- How would you communicate the probability estimates?
- How would you communicate derived inferences and decisions?
- What about inferences depending on 1000 random variables with varying reliabilities?

Unknown dependence

▶ Nature of dependence between two events A and B often not known



Unknown dependence and Fréchet's bounds

Nature of dependence between two events A and B often not known



Fréchet's bounds:

$$P(A \land B) \in \left[\max\{0, P(A) + P(B) - 1\}, \min\{P(A), P(B)\}
ight]$$

 $P(A \lor B) \in \left[\max\{P(A), P(B)\}, \min\{1, P(A) + P(B)\}
ight]$

Can probability theory differentiate between aleatoric and epistemic uncertainty?

[Aleatoric and epistemic uncertainty] must be treated *differently*;

variability should be modeled as randomness with the methods of probability theory;

incertitude should be modeled as ignorance with the methods of interval analysis.



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Interval analysis

- Computation with intervals instead of with single numbers
- Example: [2,4] [3,5] = [-3,1]
- Ideal is to obtain tightest bounds
- In practice often outer bounds are used for computational reasons

Ferson (2004)

Rational agents

In Savage's classical account of *Subjective Expected Utility Theory*, a **'rational' agent**

- models uncertainty in a problem using a single probability measure
- chooses between alternatives by maximizing expected utility

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Urn example

Urn with 20 red, 10 black, and 10 white balls The rational agent uses a pmf with p_R = 1/2 and p_B = p_W = 1/4

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Urn example

- Urn with 20 red, 10 black, and 10 white balls The rational agent uses a pmf with $p_{\rm R} = 1/2$ and $p_{\rm B} = p_{\rm W} = 1/4$
- Consider three gambles whose payoff depends on the outcome of a random draw from the urn:

	R	В	W
f _{RB}	\$100	\$100	\$0
f _{RW}	\$100	\$0	\$100
f _{BW}	\$0	\$100	\$100

The agent is *indifferent* between f_{RB} and f_{RW} , which they *strictly prefer* to f_{BW}

Risk aversion

- ▶ Urn with 20 red, 20 gray balls
- ► Two gambles:

R G f_{RG} \$50 \$50 f_R \$100 \$0

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Ambiguity aversion

- ► Two urns; the agent must choose one:
 - E 20 black balls, 20 white balls
 - ${\sf U}\,$ unknown proportion of black and white balls

One gamble:

B W f_B \$100 \$0

Risk aversion

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Ambiguity aversion

- ► Two urns; the agent must choose one:
 - E 20 black balls, 20 white balls
 - ${\sf U}\,$ unknown proportion of black and white balls
- One gamble:
- B W f_B \$100 \$0
- Rational agents must choose a pmf for U; in case they choose p_B = p_W = 1/2 they are *indifferent* between urns E and U
- Real agents strictly prefer urn E over urn U

Ellsberg paradox

 Urn with 20 red balls and 40 white or black balls in unknown proportion

► Four gambles:

	R	В	W
f _R	\$100	\$0	\$0
f _B	\$0	\$100	\$0
$f_{\rm RW} = f_{\rm R} + f_{\rm W}$	\$100	\$0	\$100
$f_{\rm BW} = f_{\rm B} + f_{\rm W}$	\$0	\$100	\$100

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• Choose between f_R and f_B

• Choose between f_{RW} and f_{BW}

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- Choose between f_R and f_B
 Real agents
 strictly prefer f_R over f_B,
 for a rational agent implying
 p_R > p_B
- Choose between f_{RW} and f_{BW}
 Real agents
 strictly prefer f_{BW} over f_{RW},
 for a rational agent implying
 p_R < p_B

Why go beyond probability theory?

To be able to deal with epistemic uncertainty:

- Distinguish sample size in uncertainty representation
- Express partial or missing knowledge

To let agents act according to a less restrictive definition of rationality:

- Be able to reflect justified aversions
- Model behavior that would otherwise be paradoxical

Overview

Kick-off (slot 1)

Classical probability theory (slot 1)

Interpretation of probability (slot 2)

Limitations of probability theory (slot 2)

Probability intervals (slot 3) Illustration: Blood groups Representation Reasoning

Credal sets (slot 3-4)

Acceptability & Desirability (slot 4–5)

Interval expectation & probability (slot 5-6)

Problem setup

- ► A partial information sample:
 - A,B AB O A,B A,B A,B A,B O
- A disease
- Two treatments with differing effectiveness:

	А	В	AB	0
f_1	0.5	0.6	0.7	0.1
f_2	0.4	0.3	0.3	0.8

Representation

PMF from observed frequencies:

p_{A}	p_{B}	p_{AB}	$p_{\rm O}$
2	2	1	2
1.0	1	8	8

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Representation

PMF from observed frequencies:

p_{A}	p_{B}	p_{AB}	$p_{\rm O}$
$[0, \frac{5}{8}]$	$[0, \frac{5}{8}]$	$\frac{1}{8}$	$\frac{2}{8}$

Reasoning

- Bounds on P({A, O})?
- Bounds on $E_p(f_1)$?
- Outcome with maximal lower probability?
- Treatment with highest upper expected effectiveness?

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Representation

PMF from observed frequencies:

p_{A}	p_{B}	p_{AB}	$p_{\rm O}$
$[0, \frac{5}{8}]$	$[0, \frac{5}{8}]$	$\frac{1}{8}$	$\frac{2}{8}$

Reasoning

- ▶ $P(\{A, O\}) \in [\frac{2}{8}, \frac{7}{8}]$
- $E_{\rho}(f_1) \in = \left[\frac{3.4}{8}, \frac{3.9}{8}\right] = \left[0.425, 0.4875\right]$
- Outcome O has maximal lower probability
- Both treatments have equal upper expected effectiveness

Basic setup, axioms, and terminology

Basic setup of the theory of probability intervals (Campos, Huete, Moral, 1994):

- Random variable X
- Finite set of outcomes \mathcal{X}
- ▶ Each outcome $x \in \mathcal{X}$ is assigned lower and upper probability values (p_x, \overline{p}_x)

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Axioms

A **probability interval** $(\underline{p}, \overline{p})$, a pair of lower and upper probability mass functions, must be:

- 1. ?
- 2. ?
- 3. ?

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A **probability interval** $(\underline{p}, \overline{p})$, a pair of lower and upper probability mass functions, must be:

- 1. Bounded: $0 \leq \underline{p}_x \leq \overline{p}_x \leq 1$ for all outcomes $x \in \mathcal{X}$
- 2. Proper: $\sum_{x \in \mathcal{X}} \underline{p}_x \leq 1 \leq \sum_{x \in \mathcal{X}} \overline{p}_x$

3. **?**




Unreachable bounds require tightening



A B O p 3/8 3/8 p 1/8 1/8

Unreachable bounds require tightening



A B O p 3/8 3/8 7/8 p 1/8 1/8 1/8





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A B O p 3/8 3/8 6/8 p 1/8 1/8 2/8

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- 3. Reachable: $\underline{p}_x \ge 1 \sum_{z \neq x} \overline{p}_z$ and $\overline{p}_x \le 1 \sum_{z \neq x} \underline{p}_z$ for all outcomes $x \in \mathcal{X}$

 A probability interval (p, p̄) is included in or less imprecise than a probability interval (q, q̄) if for all x ∈ X

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 The corresponding credal sets will also respect the same inclusion relationship





Lower and upper probability mass sums

$$P_{\widetilde{Z}}(S) := \sum_{x \in S} \underline{p}_x, \qquad \widetilde{P}(S) := \sum_{x \in S} \overline{p}_x$$

Lower probability

$$\underline{P}(S) := \max\{\underline{P}(S), 1 - \widetilde{P}(S^{c})\}$$

Upper probability

$$\overline{P}(S) := \min\{\widetilde{P}(S), 1 - \mathop{P}_{\sim}(S^{\mathsf{c}})\}$$

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$$\begin{array}{ccccc} A & B & O & AB \\ \overline{p} & 3/8 & 3/8 & 5/8 & 5/8 \\ \underline{p} & 1/8 & 1/8 & 3/8 & 0/8 \end{array}$$

Inferences

- ► Lower probability <u>P</u>({A,B})?
- Upper probability $\overline{P}(\{B,O\})$?

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Inferences

• $\underline{P}(\{A, B\}) = 2/8$ • $\overline{P}(\{B, O\}) = 7/8$

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 $[\underline{p}_x, \overline{p}_x] \subseteq [\underline{q}_x, \overline{q}_x]$

- The corresponding credal sets will also respect the same inclusion relationship
- Probability bounds will then respect the same inclusion relationship for all S ⊆ X:

 $[\underline{P}(S), \overline{P}(S)] \subseteq [\underline{Q}(S), \overline{Q}(S)]$





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Credal sets (slot 3–4) Representation Reasoning Multivariate credal sets

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Set of pmfs (probability measures) determined by *non-strict constraints* specified as

> a probability interval



Set of pmfs (probability measures) determined by *non-strict constraints* specified as

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- interval (lower & upper) probabilities or expectations (later)



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Directly specified credal sets

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discrete sets ('sets of Bayesians')



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- neighborhoods of specific pmfs ('robust Bayesians')



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Convex and closed sets Computationally convenient

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Set of pmfs (probability measures) specified *directly and explicitly* as such by

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Generally non-convex and not closed Can have impact on, e.g., decision rules

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- Set of outcomes X
- ► A credal set of probability mass functions for X is specified

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Axioms

A credal set \mathcal{C} must be:

- 1. A subset of the set of all probability mass functions for $X: \mathcal{C} \subseteq \mathcal{P}_X$
- 2. Non-empty: $\mathcal{C} \neq \emptyset$

Special types of credal sets

The vacuous credal set

$$\mathcal{P}_X^S := \{p \in \mathcal{P}_X : P(S) = 1\}$$

relative to some event $S \subseteq \mathcal{X}$ expresses that $X \in S$ and nothing more

- The vacuous credal set \$\mathcal{P}_X^X = \mathcal{P}_X\$ expresses complete ignorance
- Singleton credal sets {p} correspond to the unique pmf p they contain
- Linear-vacuous or ε-contamination credal sets are a simple neighborhood model:

$$\mathcal{C}^{oldsymbol{p},arepsilon} := ig\{(1-arepsilon)oldsymbol{p} + arepsilonoldsymbol{q}: oldsymbol{q} \in \mathcal{P}_Xig\}$$





Lower & upper probability as lower & upper envelopes

$$\underline{P}(S) := \inf_{p \in \mathcal{C}} P_p(S) \qquad \overline{P}(S) := \sup_{p \in \mathcal{C}} P_p(S)$$

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Envelope calculation as linear optimization over the credal set

$$\underline{E}(f) = \inf_{p \in \mathcal{C}} E_p(f) = \inf_{p \in \mathcal{C}} \sum_{x \in \mathcal{X}} p_x f(x)$$

For finite or closed credal sets \mathcal{C} , inf/sup becomes min/max

- ► linear-vacuous credal set $C^{p,\varepsilon}$
- mixture coefficient $\varepsilon = \frac{1}{3}$
- ▶ pmf

$$p = (p_{A}, p_{B}, p_{O}) = (\frac{1}{4}, \frac{1}{4}, \frac{2}{4})$$



▶ function
f = (f(A), f(B), f(O)) = (1, 0, -1)

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Inferences

Lower probability



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Lower probability

 $\underline{\textit{P}}(\{\mathsf{B},\mathsf{O}\})$



• function f = (f(A), f(B), f(O)) = (1, 0, -1)

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▶ function
f = (f(A), f(B), f(O)) = (1, 0, -1)

Inferences

Lower probability

$$\underline{P}(\{\mathsf{B},\mathsf{O}\}) = \min_{q \in \mathcal{P}_X} \left(\frac{2}{3}(p_{\mathsf{B}} + p_{\mathsf{O}}) + \frac{1}{3}(q_{\mathsf{B}} + q_{\mathsf{O}}) \right)$$

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Lower expectation

 $\underline{E}(f)$

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Inferences

Ε

Lower probability

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$$= \frac{2}{3}\frac{3}{4} + \frac{1}{3}0$$
$$= \frac{1}{2}$$

$$\begin{aligned} (f) &= \min_{q \in \mathcal{P}_X} \left(\frac{2}{3} \sum_{x \in \mathcal{X}} p_x f(x) + \frac{1}{3} \sum_{x \in \mathcal{X}} q_x f(x) \right) \\ &= \min_{q \in \mathcal{P}_X} \left(\frac{2}{3} (p_{\mathsf{A}} - p_{\mathsf{O}}) + \frac{1}{3} (q_{\mathsf{A}} - q_{\mathsf{O}}) \right) \end{aligned}$$
- ► linear-vacuous credal set C^{p,ε}
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Lower expectation

$$\underline{\underline{E}}(f) = \min_{q \in \mathcal{P}_X} \left(\frac{2}{3} \sum_{x \in \mathcal{X}} p_x f(x) + \frac{1}{3} \sum_{x \in \mathcal{X}} q_x f(x) \right)$$
$$= \min_{q \in \mathcal{P}_X} \left(\frac{2}{3} (p_A - p_O) + \frac{1}{3} (q_A - q_O) \right)$$
$$= \frac{2}{3} (-\frac{1}{4}) + \frac{1}{3} (-1)$$
$$= -\frac{1}{2}$$

Lower & upper probability as lower & upper envelopes

$$\underline{P}(S) := \inf_{p \in \mathcal{C}} P_p(S) \qquad \overline{P}(S) := \sup_{p \in \mathcal{C}} P_p(S)$$

Lower & upper expectation as lower & upper envelopes

$$\underline{E}(f) := \inf_{p \in \mathcal{C}} E_p(f) \qquad \overline{E}(f) := \sup_{p \in \mathcal{C}} E_p(f)$$

Envelope calculation as linear optimization over the credal set's extreme points

$$\underline{E}(f) = \inf_{p \in \mathcal{C}} E_p(f) = \inf_{p \in \mathcal{C}} \sum_{x \in \mathcal{X}} p_x f(x) = \inf_{p \in \text{ext} \mathcal{C}} \sum_{x \in \mathcal{X}} p_x f(x)$$

(Extreme points exist for closed credal sets and fully characterize convex ones.)

- ► linear-vacuous credal set $C^{p,\varepsilon}$
- mixture coefficient $\varepsilon = \frac{1}{3}$
- ▶ pmf

$$p = (p_{A}, p_{B}, p_{O}) = (\frac{1}{4}, \frac{1}{4}, \frac{2}{4})$$



▶ function
f = (f(A), f(B), f(O)) = (1, 0, -1)

Inferences

Lower probability



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<u>*P*(</u>{B,O})



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Inferences

Lower probability

$$\underline{P}(\{\mathsf{B},\mathsf{O}\}) = \min_{q \in \mathsf{ext}\,\mathcal{C}^{p,\varepsilon}}(q_\mathsf{B} + q_\mathsf{O})$$

► Lower expectation

- ► linear-vacuous credal set $C^{p,\varepsilon}$
- mixture coefficient $\varepsilon = \frac{1}{3}$
- ▶ pmf

$$p = (p_{A}, p_{B}, p_{O}) = (\frac{1}{4}, \frac{1}{4}, \frac{2}{4})$$





Inferences

Lower probability

$$\underline{P}(\{B,O\}) = \min_{q \in ext C^{p,\varepsilon}} (q_B + q_O)$$

= min $\{\frac{3}{6}, \frac{5}{6}, \frac{5}{6}\}$
= $\frac{3}{6} = \frac{1}{2}$

Lower expectation

- ► linear-vacuous credal set $C^{p,\varepsilon}$
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$$\underline{E}(f) = \min_{q \in \mathsf{ext} \, \mathcal{C}^{p,\varepsilon}} \sum_{x \in \mathcal{X}} q_x f(x)$$
$$= \min_{q \in \mathsf{ext} \, \mathcal{C}^{p,\varepsilon}} (q_A - q_O)$$

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Inferences

Lower probability

$$\underline{P}(\{\mathsf{B},\mathsf{O}\}) = \min_{q \in \mathsf{ext}\,\mathcal{C}^{p,\varepsilon}}(q_{\mathsf{B}} + q_{\mathsf{O}})$$
$$= \min\{\frac{3}{6}, \frac{5}{6}, \frac{5}{6}\}$$
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Lower expectation

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$$= \min\{\frac{1}{6}, -\frac{1}{6}, -\frac{3}{6}\}$$
$$= -\frac{3}{6} = -\frac{1}{2}$$

Multivariate credal sets: basic setup & basic idea

• Index set
$$N = \{1, \ldots, n\}$$

- Multivariate variable $\boldsymbol{X} = (X_1, X_2, \dots, X_n)$
- ▶ Set of possible outcomes $x \in X = X_1 \times X_2 \times \ldots \times X_n$
- A joint credal set $\mathcal{C}^{\mathbf{X}}$ of joint probability mass functions for \mathbf{X} is specified

Apply probabilistic operations pointwise to the elements of the credal sets involved

Marginal credal sets

A marginal credal set is defined for a subset of the random variables

- Let $K \subseteq N$, then
 - $\blacktriangleright \ \pmb{X}_{\mathcal{K}} := (X_k : k \in \mathcal{K}) \text{ and } \pmb{x}_{\mathcal{K}} \in \mathcal{X}_{\mathcal{K}} := {\color{black} \times_{k \in \mathcal{K}} \mathcal{X}_k}$
 - ▶ Notation: $p^{X_{\kappa}}$ is the X_{κ} -marginal of the joint pmf p^{X}

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A marginal credal set is obtained by marginalizing each of its member pmfs:

$$\mathcal{C}^{oldsymbol{x}_{K}}=\left\{ p^{oldsymbol{x}_{K}}:p^{oldsymbol{x}}\in\mathcal{C}^{oldsymbol{x}}
ight\}$$

Problem setup

- ► Two random variables X₁ and X₂ with outcome spaces X₁ = {0,1} and X₂ = {−,+}
- Joint credal set C^X := C^{p,ε} with ε unspecified and p given below (in black), together with its marginals (in green)

р	0	1	p^{X_2}
—	3/9	1/9	4/9
+	2/9	3/9	5/9
p^{X_1}	5/9	4/9	

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Inference

Marginal credal set \mathcal{C}^{X_1}

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Inference

Marginal credal set C^{X_1}

Solution

• Marginalize element $q = (1 - \varepsilon)p + \varepsilon r$ of $\mathcal{C}^{\mathbf{X}}$, where $r \in \mathcal{P}_{\mathbf{X}}$:

$$q_{x_1}^{X_1} = \sum_{x_2 \in \mathcal{X}_2} q_{(x_1, x_2)} = (1 - arepsilon) p_{x_1}^{X_1} + arepsilon r_{x_1}^{X_1}$$

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• Gather all marginalized elements:

$$\mathcal{C}^{X_1} = \mathcal{C}^{p^{X_1},\varepsilon}$$

because r^{X_1} ranges over \mathcal{P}_{X_1} as r ranges over $\mathcal{P}_{\boldsymbol{X}}$

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$$\begin{array}{c|c|c|c|c|c|c|c|}\hline q & 0 & 1 \\ \hline - & (1-\varepsilon)^{3/9} + \varepsilon r_{(0,-)} & (1-\varepsilon)^{1/9} + \varepsilon r_{(1,-)} \\ + & (1-\varepsilon)^{2/9} + \varepsilon r_{(0,+)} & (1-\varepsilon)^{3/9} + \varepsilon r_{(1,+)} \\ \hline q^{\chi_1} & (1-\varepsilon)^{5/9} + \varepsilon r_0^{\chi_1} & (1-\varepsilon)^{4/9} + \varepsilon r_1^{\chi_1} \\ \hline \end{array}$$

Conditional credal sets

- A conditional credal set is determined by
 - assuming some event is known to be true, or specifically
 - ▶ assuming some random variables take some given values: $X_K = x_K$, with $K \subset N$

▶ Notation: $p^{X_{N\setminus K}|x_K}$ is the X_K -conditional of the joint pmf p^X

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- A conditional credal set is obtained by conditioning each of its elements for which this operation is defined:

$$\mathcal{C}^{\boldsymbol{X}_{N\setminus K}|\boldsymbol{x}_{K}} = \left\{ p^{\boldsymbol{X}_{N\setminus K}|\boldsymbol{x}_{K}} : \left(p^{\boldsymbol{X}} \in \mathcal{C}^{\boldsymbol{X}} \land p^{\boldsymbol{X}_{K}}(\boldsymbol{x}_{K}) > 0 \right) \right\}$$

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This is called conditioning by regular extension

Problem setup

- ► Two random variables X₁ and X₂ with outcome spaces X₁ = {0, 1} and X₂ = {−, +}
- Joint credal set C^X := C^Γ with elements q given below (in black), together with its marginals (in green), with γ ∈ Γ = [-³/₉, ³/₉]

$$\begin{array}{c|c|c|c|c|c|c|c|}\hline q & 0 & 1 & q^{X_2} \\ \hline - & 3/9 - \gamma & 1/9 & 4/9 - \gamma \\ + & 2/9 & 3/9 + \gamma & 5/9 + \gamma \\ \hline q^{X_1} & 5/9 - \gamma & 4/9 + \gamma & \end{array}$$

Inference

Conditional credal set $\mathcal{C}^{X_1|+}$

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Inference

Conditional credal set $\mathcal{C}^{X_1|+}$

Solution

• Condition elements q of $\mathcal{C}^{\mathbf{X}}$:

$$q^{X_1|+} = rac{q_{(X_1,+)}}{q_+^{X_2}} = rac{(2/9,3/9+\gamma)}{5/9+\gamma} = (r,1-r),$$

with
$$r = rac{2/9}{5/9+\gamma}$$

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$$\mathcal{C}^{X_1|+}=\left\{(r,1-r):r\in [rac{1}{4},1]
ight\}$$

Problem setup

- \blacktriangleright Two random variables X_1 and X_2 with outcome spaces $\mathcal{X}_1 = \{0, 1\}$ and $X_2 = \{-,+\}$
- ▶ loint credal set $C^{\mathbf{X}} := C^{\Gamma}$ with elements q given below (in black), together with its marginals (in green), with $\gamma \in \Gamma = \left[-\frac{3}{6}, \frac{3}{6}\right]$

$$\begin{array}{c|c|c|c|c|c|c|c|c|}\hline q & 0 & 1 & q^{X_2} \\ \hline - & 3/9 - \gamma & 1/9 & 4/9 - \gamma \\ + & 2/9 & 3/9 + \gamma & 5/9 + \gamma \\ \hline q^{X_1} & 5/9 - \gamma & 4/9 + \gamma & \hline \end{array}$$

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Solution

• Condition elements q of $\mathcal{C}^{\mathbf{X}}$:

$$q^{X_1|+} = rac{q_{(X_1,+)}}{q_+^{X_2}} = rac{(2/9,3/9+\gamma)}{5/9+\gamma} = (r,1-r),$$

with
$$r=rac{2/9}{5/9+\gamma}$$



Gather all conditioned elements:

$$\mathcal{C}^{X_1|+} = \left\{ (r, 1-r) : r \in [rac{1}{4}, 1]
ight\}$$

(Exercise: show that $\mathcal{C}^{X_1|+} = \mathcal{C}^{(1,0),\frac{3}{4}}$)

Complete independence for credal sets

- Different generalizations of (stochastic) independence to credal sets are possible
- Here, we only consider the independence concept associated to the 'sets of Bayesians' interpretation
- Notation: $p \otimes q$ denotes the independent product of two pmfs p and q

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Definition of complete independence

Consider $\mathbf{X} = (X_1, X_2)$; the random variables X_1 and X_2 are *completely independent* if they are stochastically independent for each of the pmfs in the joint credal set $\mathcal{C}^{\mathbf{X}}$:

$$\mathcal{C}^{\boldsymbol{X}} \subseteq \mathcal{C}^{\boldsymbol{X}_1} \otimes \mathcal{C}^{\boldsymbol{X}_2} := \left\{ \boldsymbol{p}^{\boldsymbol{X}_1} \otimes \boldsymbol{p}^{\boldsymbol{X}_2} : \boldsymbol{p}^{\boldsymbol{X}_1} \in \mathcal{C}^{\boldsymbol{X}_1}, \boldsymbol{p}^{\boldsymbol{X}_2} \in \mathcal{C}^{\boldsymbol{X}_2}
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Complete independence for credal sets

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ight\}$$

Not convex even if C^{X_1} and C^{X_2} are!

Problem setup

- ▶ Random variables X_1 and X_2 with outcome spaces $\mathcal{X}_1 = \{0, 1\}$ and $\mathcal{X}_2 = \{-, +\}$
- Marginal credal sets $\mathcal{C}^{X_1} := \mathcal{C}^{(\frac{1}{2},\frac{1}{2}),\frac{1}{2}}$ and $\mathcal{C}^{X_2} := \mathcal{C}^{(\frac{1}{3},\frac{2}{3}),\frac{1}{3}}$

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Goal

Show that their completely independent joint credal set is not convex

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Goal

Show that their completely independent joint credal set is not convex

Demonstration

We construct a counterexample to convexity, i.e., we construct a convex mixture of elements of the joint that lies outside the joint because it does not factorize

Problem setup

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Goal

Show that their completely independent joint credal set is not convex

- We construct a counterexample to convexity, i.e., we construct a convex mixture of elements of the joint that lies outside the joint because it does not factorize
- ▶ An element q of $C^{\mathbf{X}} = C^{X_1} \otimes C^{X_2}$ is defined for every $r^{X_1} \in \mathcal{P}_{X_1}$ and $r^{X_2} \in \mathcal{P}_{X_2}$ by

$$q_{(x_1,x_2)} = (\frac{1}{2}p_{x_1}^{X_1} + \frac{1}{2}r_{x_1}^{X_1})(\frac{2}{3}p_{x_2}^{X_2} + \frac{1}{3}r_{x_2}^{X_2})$$

Problem setup

▶ Random variables X_1 and X_2 with outcome spaces $\mathcal{X}_1 = \{0, 1\}$ and $\mathcal{X}_2 = \{-, +\}$

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5/9 4/9

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Goal

Show that their completely independent joint credal set is not convex

$$r_{0}^{X_{1}} = r_{-}^{X_{2}} = 0$$

$$r_{0}^{X_{1}} = r_{-}^{X_{2}} = 1$$

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Goal

Show that their completely independent joint credal set is not convex

$$r_{0}^{X_{1}} = r_{-}^{X_{2}} = 0 \qquad r_{0}^{X_{1}} = r_{-}^{X_{2}} = 1$$

$$\frac{1}{2} \frac{\begin{vmatrix} 0 & 1 & q^{X_{2}} \\ - & \frac{2}{36} & \frac{6}{36} & \frac{2}{9} \\ + & \frac{7}{36} & \frac{21}{36} & \frac{7}{9} \\ \frac{1}{q^{X_{1}}} & \frac{1}{4} & \frac{3}{4} \end{vmatrix} + \frac{1}{2} \frac{\begin{vmatrix} 0 & 1 & q^{X_{2}} \\ - & \frac{15}{36} & \frac{5}{36} & \frac{5}{9} \\ + & \frac{12}{36} & \frac{4}{36} & \frac{4}{9} \\ \frac{1}{q^{X_{1}}} & \frac{3}{4} & \frac{1}{4} \end{vmatrix} = \frac{\begin{vmatrix} 0 & 1 & q^{X_{2}} \\ - & \frac{17}{772} & \frac{11}{72} \\ + & \frac{19}{72} & \frac{25}{72} \\ \frac{1}{q^{X_{1}}} \end{vmatrix}$$

Problem setup

▶ Random variables X_1 and X_2 with outcome spaces $\mathcal{X}_1 = \{0, 1\}$ and $\mathcal{X}_2 = \{-, +\}$

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Goal

Show that their completely independent joint credal set is not convex

$$r_{0}^{X_{1}} = r_{-}^{X_{2}} = 0 \qquad r_{0}^{X_{1}} = r_{-}^{X_{2}} = 1 \qquad \text{not factorizing}$$

$$\frac{1}{2} \frac{\begin{vmatrix} 0 & 1 & q^{X_{2}} \\ - & \frac{2}{36} & \frac{6}{36} & \frac{2}{9} \\ + & \frac{7}{36} & \frac{21}{36} & \frac{7}{9} \\ \frac{1}{q^{X_{1}}} & \frac{1}{4} & \frac{3}{4} \end{vmatrix} + \frac{1}{2} \frac{\begin{vmatrix} 0 & 1 & q^{X_{2}} \\ - & \frac{15}{36} & \frac{5}{36} & \frac{5}{9} \\ + & \frac{12}{36} & \frac{4}{36} & \frac{4}{9} \\ \frac{1}{q^{X_{1}}} & \frac{3}{4} & \frac{1}{4} \end{vmatrix} = \frac{\begin{vmatrix} 0 & 1 & q^{X_{2}} \\ - & \frac{17}{72} & \frac{17}{12} & \frac{7}{18} \\ + & \frac{19}{72} & \frac{25}{72} & \frac{11}{18} \\ \frac{1}{q^{X_{1}}} & \frac{1}{2} & \frac{1}{2} & \frac{1}{2} \end{vmatrix}$$

Overview

Kick-off (slot 1)

Classical probability theory (slot 1)

Interpretation of probability (slot 2)

Limitations of probability theory (slot 2)

Probability intervals (slot 3)

Credal sets (slot 3–4)

Acceptability & Desirability (slot 4–5) Representation Reasoning Axiom variants Multivariate acceptability

Interval expectation & probability (slot 5-6
Gambles

Earlier appearance of 'gamble':

- In the betting game, indicators 1_x and 1_S
- ln urn problems, scaled indicators $f_i \propto 1_S$
- In blood groups example, effectiveness descriptions f₁, f₂, positive functions on X

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Exchanges

- In the betting game, differences between fair prices and gambles p_x − 1_x = p_x · 1_X − 1_x and P(S) − 1_S
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Linear space of gambles

(Bounded) real-valued functions in $\mathcal{L} = (\mathcal{X} \to \mathbb{R})$, as foundation for representation (or linear subspace of functions instead)

Basic setup, axioms, and terminology

Basic setup of acceptability:

- Random variable X
- Finite set of outcomes \mathcal{X}
- ▶ Each gamble $f \in \mathcal{L}$ is acceptable $(f \in D)$ or not $(f \notin D)$

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Axioms

A set of acceptable gambles \mathcal{D} is coherent when:

D1. if $f \leq 0$, then $f \notin D$ (avoiding sure loss)

D2. if $f \ge 0$, then $f \in D$ (accepting nonnegative gain)

D3. if $f \in \mathcal{D}$ and $\lambda \in \mathbb{R}_{>}$, then $\lambda \cdot f \in \mathcal{D}$ (positive scaling)

D4. if $f, g \in D$, then $f + g \in D$ (combination)

Gamble (vector) inequalities: $f \ge 0$ iff min $f \ge 0$, $f \ge 0$ iff min $f \ge 0$, f > 0 iff $f \ge 0$ and $f \ne 0$

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A set of acceptable gambles ${\mathcal D}$ is coherent when:

- D1. $\mathcal{L}_{\sphericalangle} \cap \mathcal{D} = \emptyset$ (avoiding sure loss)
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- D3. $\mathbb{R}_{>} \cdot \mathcal{D} \subseteq \mathcal{D}$ (positive scaling)
- D4. $\mathcal{D} + \mathcal{D} \subseteq \mathcal{D}$ (combination)

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Visualizing sets of acceptable gambles

The avoiding loss and accepting gain axioms constrain $\ensuremath{\mathcal{D}}$



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Visualizing sets of acceptable gambles

The avoiding loss and accepting gain axioms constrain $\ensuremath{\mathcal{D}}$

Coherent sets \mathcal{D} are *cones* because of positive scaling and combination



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 f and g
- Depending on where their differences
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Background preference/order implied by D1 and D2 $\,$

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An $\textbf{assessment}~\mathcal{A}$ is a set of gambles the subject finds acceptable

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Directly specified



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The natural extension *E(A)* avoids sure loss (satisfies D1) iff posi *A* avoids sure loss



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Set of all coherent sets of acceptable gambles

- ► This set can be (partially) ordered according to set inclusion ⊂
- ▶ If $\mathcal{D}_1 \subset \mathcal{D}_2$, then \mathcal{D}_1 is called less committal than \mathcal{D}_2
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Coherent extensions of an assessment ${\mathcal A}$ that avoids sure loss

- \blacktriangleright In general, there are multiple coherent extensions of ${\cal A}$
- The least committal extension is the smallest one: it adds the least commitments (it is *conservative* in that regard)
- ▶ The least committal extension is equal to the intersection of all coherent extensions
- The least committal extension coincides with the natural extension



Changing how \mathcal{D} is constrained

- ▶ Replacing the background sets $\mathcal{L}_{\leq}/\mathcal{L}_{\geq}$ in Axiom D1/D2 (extra relevant for infinite \mathcal{X})
- Adding additional axioms constraining, e.g., topological closure

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- *L* ≤ /*L*≥ and *D* closed

 (almost desirability, Walley, 1991)



L < /*L*≥ (favorable gambles, Seidenfeld et al., 1990; Walley, 2000)



 L_≤/L_> (desirability, De Cooman, Thursday! D1 becomes 0 ∉ D)





Changing what shape ${\mathcal D}$ can take

- Replacing the generating rules in Axioms D3 and D4
- Adding additional generating rules

Changing what shape \mathcal{D} can take

- Replacing the generating rules in Axioms D3 and D4
- Adding additional generating rules
- Example: replace positive scaling and combination by convexity,

 $\begin{array}{l} \text{if } f,g\in\mathcal{D} \text{ and } \mu\in[0,1],\\ \text{then } \mu\cdot f+(1-\mu)\cdot g\in\mathcal{D} \end{array}$

This means replacing the *positive linear hull* posi by the *convex hull* co



Multivariate acceptability: basic setup

- Index set $N = \{1, \ldots, n\}$
- Multivariate variable $\boldsymbol{X} = (X_1, X_2, \dots, X_n)$
- ▶ Set of possible outcomes $\mathbf{x} \in \mathcal{X} = \mathcal{X}_1 \times \mathcal{X}_2 \times \ldots \times \mathcal{X}_n$
- Linear space of gambles \mathcal{L} on \mathcal{X}
- \blacktriangleright We assume given a coherent set of acceptable gambles $\mathcal{D} \subset \mathcal{L}$

Marginal set of acceptable gambles

- \blacktriangleright A marginal set of acceptable gambles is defined on a linear subspace of ${\cal L}$
- ▶ We consider subspaces defined by $\mathcal{L}_{\mathcal{K}} := (\mathcal{X}_{\mathcal{K}} \to \mathbb{R})$, with $\mathcal{K} \subseteq \mathcal{N}$
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- ► For this, we lift it to its cylindrical extension $\uparrow_{X_{N\setminus K}} f$ defined by

$$(\uparrow_{\boldsymbol{X}_{N\setminus K}} f)(\boldsymbol{x}) = f(\boldsymbol{x}_K)$$

and check whether that is acceptable

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The marginal set of acceptable gambles is the inverse image of the cylindrical extension:

$$\mathcal{D}_{\mathcal{K}} := \uparrow_{oldsymbol{X}_{N\setminus \mathcal{K}}}^{-1} \mathcal{D} = \left\{ f \in \mathcal{L}_{\mathcal{K}} : \ \uparrow_{oldsymbol{X}_{N\setminus \mathcal{K}}} f \in \mathcal{D}
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Conditional set of acceptable gambles

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 - assuming some event is known to be true, or specifically
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$$(\uparrow \boldsymbol{x}_{\kappa} f) \cdot 1 \boldsymbol{x}_{\kappa} = \boldsymbol{x}_{\kappa}$$

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and check whether that is acceptable

▶ The conditional set of acceptable gambles is the inverse image of this procedure:

$$\mathcal{D} \sqcup (oldsymbol{X}_{K} = oldsymbol{x}_{K}) := \left\{ f \in \mathcal{L}_{N \setminus K} : (\uparrow_{oldsymbol{X}_{K}} f) \cdot 1_{oldsymbol{X}_{K} = oldsymbol{x}_{K}} \in \mathcal{D}
ight\}$$

Epistemic irrelevance & independence

- We here consider the case |N| = 2
- ► The random variable X₂ is epistemically irrelevant to X₁ if the X₂-conditionals coincide with the X₁-marginal for all x₂ ∈ X₂:

$$\mathcal{D} \rfloor (X_2 = x_2) = \mathcal{D}_1$$

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- Structural assumptions such as these can be combined with natural extension

Overview

Kick-off (slot 1)

Classical probability theory (slot 1)

Interpretation of probability (slot 2)

Limitations of probability theory (slot 2)

Probability intervals (slot 3)

Credal sets (slot 3–4)

Acceptability & Desirability (slot 4–5)

Interval expectation & probability (slot 5–6) Representation Reasoning Representation (continued) Learning Multivariate lower expectations

Probability interval example



Interval probability example





Interval expectation example



Basic setup, axioms, and terminology

Basic setup of the theory of coherent interval probability & expectation:

- Random variable X
- Set of outcomes \mathcal{X}
- \blacktriangleright Set of events $\mathcal{S}\subseteq 2^{\mathcal{X}}$ or set of gambles $\mathcal{F}\subseteq \mathcal{L}$

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The lower and upper expectation operators \underline{E} and \overline{E} , must be:

7

Eliciting lower & upper expectations with a relaxed betting game

- ► Two players:
 - subject: gambler, states acceptable exchanges, their 'assessment'
 - bookie: chooses combination of acceptable exchanges
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$$\begin{array}{c|c} \text{exchange } f - \alpha & \text{exchange } \beta - f \\ \hline \text{acceptable} & \text{indeterminate} & \text{acceptable} \\ \hline \alpha & \underline{E}(f) & \overline{E}(f) & \beta \end{array}$$

- The subject states $\underline{E}(f)$ for f and $\overline{E}(g)$ for g
- ► The corresponding *assessment* is

$$\mathcal{A} = \left\{ f - \alpha \cdot \mathbf{1}_{\mathcal{X}} : \alpha < \underline{E}(f) \right\}$$
$$\cup$$
$$\left\{ \beta \cdot \mathbf{1}_{\mathcal{X}} - g : \beta > \overline{E}(g) \right\}$$

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Assessment as a set of acceptable gambles & marginal gambles

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• $f - \underline{E}(f)$ and $\overline{E}(g) - g$ are called marginal gambles; collect them in a set \mathcal{M} , then

$$\mathcal{A} = igcup_{h \in \mathcal{M}} \{h + arepsilon \cdot 1_{\mathcal{X}} : arepsilon \in \mathbb{R}_{>}\} = \mathcal{M} + \mathbb{R}_{>} \cdot 1_{\mathcal{X}}$$



Avoiding sure loss & natural extension

We work in the language of acceptable gambles and later translate to the language of interval expectation



Avoiding sure loss & natural extension

- We work in the language of acceptable gambles and later translate to the language of interval expectation
- ▶ posi *A* avoids sure loss iff

$$\begin{aligned} \mathsf{posi}\,\mathcal{A} \cap \mathcal{L}_{<} \neq \emptyset \Leftrightarrow \mathsf{posi}(\mathcal{M} + \mathbb{R}_{>} \cdot \mathbf{1}_{\mathcal{X}}) \cap \mathcal{L}_{<} \neq \emptyset \\ \Leftrightarrow (\mathsf{posi}\,\mathcal{M} + \mathbb{R}_{>} \cdot \mathbf{1}_{\mathcal{X}}) \cap \mathcal{L}_{<} \neq \emptyset \\ \Leftrightarrow \mathsf{posi}\,\mathcal{M} \cap \mathcal{L}_{<} \neq \emptyset \end{aligned}$$



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The natural extension is then

$$egin{aligned} \mathcal{E}(\mathcal{A}) &= (\mathsf{posi}\,\mathcal{A} + \mathcal{L}_{\geq}) \cup \mathcal{L}_{\geq} \ &= (\mathsf{posi}\,\mathcal{M} + \mathbb{R}_{>} \cdot 1_{\mathcal{X}} + \mathcal{L}_{\geq}) \cup \mathcal{L}_{\geq} \ &= (\mathsf{posi}\,\mathcal{M} + \mathcal{L}_{>}) \cup \mathcal{L}_{\geq} \end{aligned}$$



From acceptable gambles to lower & upper expectations

- Consider a coherent set of acceptable gambles D and a gamble f
- We infer lower and upper expectations for f using the betting game interpretation



From acceptable gambles to lower & upper expectations

- Consider a coherent set of acceptable gambles D and a gamble f
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 $\underline{E}(f) := \sup\{\alpha \in \mathbb{R} : f - \alpha \in \mathcal{D}\}\$



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Upper expectation

$$\overline{E}(f) := \inf\{eta \in \mathbb{R} : eta - f \in \mathcal{D}\}$$



• Consider marginal gambles \mathcal{M} , the corresponding \mathcal{A} , and some gamble f

$$\underline{E}_*(f) := \sup\{\alpha \in \mathbb{R} : f - \alpha \in \mathcal{E}(\mathcal{A})\}$$

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Upper expectation

 $\overline{E}_*(f) = \inf_{h \in \text{posi } \mathcal{M} \cup \{0\}} \max(f + h) \quad (\text{linear optimization problem over a cone})$

Natural extension illustration



Natural extension illustration



Natural extension illustration



Basic setup of the theory of coherent interval expectations:

- ▶ Random variable X ▶ Set of functions $\mathcal{F} \subseteq (\mathcal{X} \to \mathbb{R})$
- Set of outcomes \mathcal{X}
- ► Each f ∈ F is assigned a lower or upper expectation value, leading to a set of marginal gambles M

Axioms

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 $\inf_{h \in \text{posi } \mathcal{M} \cup \{0\}} \max(h - (f - \underline{E}(f))) \ge 0$

Special types of lower expectations

- The vacuous lower expectation <u>E</u>^S relative to some event S ⊆ X is defined by <u>E</u>^S(f) := min_{x∈S} f(x) and expresses that X ∈ S and nothing more
- The vacuous lower expectation <u>E</u>^X expresses complete ignorance
- The linear expectation E_p corresponding to some pmf p is also a coherent lower expectation
- ▶ The linear-vacuous lower expectation $\underline{E}_{p,\varepsilon}$ is as well:

$$\underline{E}_{\rho,\varepsilon}(f) := (1-\varepsilon)E_{\rho}(f) + \varepsilon \min_{x \in \mathcal{X}} f(x)$$



Some properties of coherent lower and upper expectations

Conjugacy
$$\overline{E}(f) = -\underline{E}(-f)$$

Boundedness $\min_{x \in \mathcal{X}} f(x) \leq \underline{E}(f) \leq \overline{E}(f) \leq \max_{x \in \mathcal{X}} f(x)$

Positive homogeneity $\underline{E}(\lambda f) = \lambda \underline{E}(f)$ for all $\lambda > 0$

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Monotonicity if $f \leq g$ then $\underline{E}(f) \leq \underline{E}(g)$ Constant additivity $\underline{E}(f + \mu \mathbf{1}_{\mathcal{X}}) = \underline{E}(f) + \mu$ for all $\mu \in \mathbb{R}$ Mixed super/sub-linearity

$$\underline{E}(f) + \underline{E}(g) \leq \underline{E}(f+g) \leq \underline{E}(f) + \overline{E}(g) \leq \overline{E}(f+g) \leq \overline{E}(f) + \overline{E}(g)$$

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Convex mixtures if \underline{E}^1 and \underline{E}^2 are coherent, then so is $\varepsilon \underline{E}^1 + (1 - \varepsilon) \underline{E}^2$ for all $0 \le \varepsilon \le 1$

Lower expectations on linear function spaces

If \mathcal{F} is a linear space $(f, g \in \mathcal{F}$ then $af + bg \in \mathcal{F}$ for all $a, b \in \mathbb{R}$) then the axioms simplify:

Boundedness $\underline{E}(f) \ge \min_{x \in \mathcal{X}} f(x)$ for all $f \in \mathcal{F}$

Positive homogeneity $\underline{E}(\lambda f) = \lambda \underline{E}(f)$ for all $f \in \mathcal{F}$ and $\lambda > 0$

Super-linearity $\underline{E}(f+g) \geq \underline{E}(f) + \underline{E}(g)$ for all $f, g \in \mathcal{F}$

Lower envelopes

- Let Γ be some index set
- ▶ Consider a set $\{\underline{E}^{\gamma} : \gamma \in \Gamma\}$ of coherent lower expectations on \mathcal{F}
- ▶ Then its lower envelope \underline{E} defined by $\underline{E}(f) := \inf_{\gamma \in \Gamma} \underline{E}^{\gamma}(f)$ on \mathcal{F} is coherent

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Lower envelopes of credal sets

- \blacktriangleright Let ${\mathcal C}$ be a non-empty credal set
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Credal sets from lower expectations

- Let \underline{E} be a lower expectation on \mathcal{F} avoiding sure loss
- ▶ Then the credal set $C^{\underline{E}} := \{p \in \mathcal{P}_X : E_p(f) \ge \underline{E}(f) \text{ for all } f \in \mathcal{F}\}$ is non-empty

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The categorical prediction task What is the color of the next ball to be drawn from an urn? The categorical prediction task

What is the color of the next ball to be drawn from an urn?

Setup

- ▶ Outcome space X (of ball colors: Red, Green, Black, White,...)
- We have n ∈ N observations x = (x₁,...,x_n) (sequence of colors of previous balls drawn)
- Draw inferences for or make decisions related to the next observation X_{n+1} (color of next ball drawn)

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Exchangeability assumption

The order of the observations is irrelevant (inferences should remain the same when (R, G, R) or (R, R, G) has been observed)

• Observed occurrence vector $\mathbf{n} \in \mathbb{N}^{\mathcal{X}}$ with $n_z := \sum_{k=1}^n \delta_{zx_k}$ and $n_S := \sum_{x \in S} n_x$ (if $\mathbf{x} = (\mathbb{R}, \mathbb{G}, \mathbb{R}, \mathbb{W})$ then $n_{\mathbb{R}} = 2$, $n_{\mathbb{G}} = 1$, $n_{\mathbb{W}} = 1$, and $n_{\neg \mathbb{W}} = 3$)

Idea

Add epistemic uncertainty using s > 0 pseudo-observations (of unknown color) and predict according to the 'observed' frequency

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Predictive inference

• Outcome prediction:
$$\underline{P}_s(\{x\}|\mathbf{n}) = \frac{n_x}{s+n}$$
 and $\overline{P}_s(\{x\}|\mathbf{n}) = \frac{s+n_x}{s+n}$

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$$\underline{E}_s(f|\boldsymbol{n}) = rac{n}{s+n} E_p(f) + rac{s}{s+n} \min_{x \in \mathcal{X}} f(x)$$
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Properties of $\underline{E}_{s}(\cdot|\mathbf{n})$

- linear-vacuous model; vacuous for n = 0; more precise with more observations
- does not depend on a specific categorization \mathcal{X}
- immediate prediction model of the imprecise Dirichlet-multinomial model ID(M)M

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Example setup

Inferences

• $\underline{P}_2(\{B, W\}|n) = ?$ • $\overline{P}_2(\{B, W\}|n) = ?$

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Example setup

Inferences

• $\underline{P}_2(\{B, W\}|\mathbf{n}) = 2/7$ • $\overline{P}_2(\{B, W\}|\mathbf{n}) = 4/7$

Multivariate interval expectation: basic setup & basic idea

• Index set
$$N = \{1, \ldots, n\}$$

- Multivariate variable $\boldsymbol{X} = (X_1, X_2, \dots, X_n)$
- ▶ Set of possible outcomes $\mathbf{x} \in \mathcal{X} = \mathcal{X}_1 \times \mathcal{X}_2 \times \ldots \times \mathcal{X}_n$
- ▶ A joint lower expectation \underline{E}^{X} is specified on some set of functions $\mathcal{F} \subseteq \mathcal{L}$

Formulate an appropriate function on \mathcal{L} whose lower expectation provides the desired inference; calculate the lower expectation using natural extension

Marginal lower expectations

A marginal lower expectation is defined for a subset of the random variables

- ▶ Let $K \subseteq N$, then
 - $\blacktriangleright \ \boldsymbol{X}_{K} := (X_{k} : k \in K) \text{ and } \boldsymbol{x}_{K} \in \mathcal{X}_{K} := \bigotimes_{k \in K} \mathcal{X}_{k}$
 - ▶ Notation: $\underline{E}^{X_{\kappa}}$ is the X_{κ} -marginal of the joint lower expectation \underline{E}^{X}

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 - ▶ Notation: $\underline{E}^{X_{\kappa}}$ is the X_{κ} -marginal of the joint lower expectation \underline{E}^{X}
- ► The marginal lower expectation for f ∈ L_K is obtained by calculating the joint natural extension of its *cylindrical extension*:

$$\underline{E}^{\boldsymbol{X}_{K}}(f) := \underline{E}_{*}^{\boldsymbol{X}}(\uparrow_{\boldsymbol{X}_{N\setminus K}} f) \quad \text{where} \quad (\uparrow_{\boldsymbol{X}_{N\setminus K}} f)(\boldsymbol{x}) \coloneqq f(\boldsymbol{x}_{K})$$

Marginal lower expectation example

Example setup

- ► Two random variables X₁ and X₂ with outcome spaces X₁ = {0, 1} and X₂ = {−, +}
- Joint lower expectation <u>E</u>^X := <u>E</u>_{p,ε} with ε unspecified and p given on the right (in black), together with its marginals (in green)

Inference

Marginal lower expectation \underline{E}^{X_1}

р	0	1	p^{X_2}
_	3/9	1/9	4/9
+	2/9	3/9	5/9
p^{X_1}	5/9	4/9	

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Inference

Marginal lower expectation \underline{E}^{X_1}

Solution

$$\underline{E}_{\rho,\varepsilon}(\uparrow_{X_2} f) = (1 - \varepsilon)E_p(\uparrow_{X_2} f) + \varepsilon \min_{x \in \mathcal{X}}(\uparrow_{X_2} f)(x_1, x_2)$$
$$= (1 - \varepsilon)E_p x_1(f) + \varepsilon \min_{x_1 \in \mathcal{X}_1} f(x_1)$$
$$= \underline{E}_p x_{1,\varepsilon}(f)$$

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- A conditional lower expectation is determined by
 - assuming some event is known to be true, or specifically
 - ▶ assuming some random variables take some given values: $X_K = x_K$, with $K \subset N$

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- ► The conditional lower expectation for f ∈ L_{N\K} is the maximum acceptable buying price for the corresponding called-off gamble:

$$\underline{E}^{\boldsymbol{X}_{N\setminus K}}(f|\boldsymbol{X}_{K}=\boldsymbol{x}_{K}):= \max\Bigl\{\alpha\in\mathbb{R}:\underline{E}_{*}^{\boldsymbol{X}}(\uparrow_{\boldsymbol{X}_{K}}(f-\alpha\cdot\boldsymbol{1}_{\mathcal{X}_{N\setminus K}})\cdot\boldsymbol{1}_{\boldsymbol{X}_{K}=\boldsymbol{x}_{K}})\geq 0\Bigr\}$$

$$\text{if }\underline{P}(\boldsymbol{X}_{K}=\boldsymbol{x}_{K})>0\text{, otherwise }\underline{E}^{\boldsymbol{X}_{N\setminus K}}(\cdot|\boldsymbol{X}_{K}=\boldsymbol{x}_{K}):=\underline{E}^{\mathcal{X}_{N\setminus K}}$$

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$$\text{if } \underline{P}(\boldsymbol{X}_{K} = \boldsymbol{x}_{K}) > 0, \text{ otherwise } \underline{E}^{\boldsymbol{X}_{N \setminus K}}(\cdot | \boldsymbol{X}_{K} = \boldsymbol{x}_{K}) := \underline{E}^{\mathcal{X}_{N \setminus K}}$$

This is called conditioning by natural extension or the generalized Bayes's rule

Example setup

- ► Two random variables X₁ and X₂ with outcome spaces X₁ = {0, 1} and X₂ = {−, +}
- ▶ Joint lower expectation <u>E</u>^X := <u>E</u>_{p,ε} with ε unspecified and p given on the right (in black), together with its marginals (in green)



Inference

Conditional lower expectation $\underline{E}^{X_1}(\cdot|X_2=+)$

Solution (start)

$$\underline{E}^{X_1}(f|X_2=+) = \max\Big\{\alpha \in \mathbb{R} : \underline{E}_{p,\varepsilon}(\uparrow_{X_2} (f - \alpha \cdot \mathbf{1}_{\mathcal{X}_1}) \cdot \mathbf{1}_{X_2=+}) \ge 0\Big\}$$

$$\underline{E}^{X_1}(f|X_2=+) = \max \Big\{ \alpha \in \mathbb{R} : \underline{E}_{p,\varepsilon}(\uparrow_{X_2} (f - \alpha \cdot \mathbf{1}_{\mathcal{X}_1}) \cdot \mathbf{1}_{X_2=+}) \geq 0 \Big\}$$

$$\underline{E}^{X_1}(f|X_2=+) = \max \Big\{ \alpha \in \mathbb{R} : \underline{E}_{\boldsymbol{\rho},\varepsilon} (\uparrow_{X_2} (f - \alpha \cdot \mathbf{1}_{\mathcal{X}_1}) \cdot \mathbf{1}_{X_2=+}) \geq 0 \Big\}$$

$$= \max \Big\{ \alpha \in \mathbb{R} : (1 - \varepsilon) E_{\rho}(\uparrow_{X_2} (f - \alpha \cdot \mathbf{1}_{\mathcal{X}_1}) \cdot \mathbf{1}_{X_2 = +}) + \varepsilon \min_{\mathbf{x} \in \mathcal{X}} (f(\mathbf{x}_1) - \alpha) \cdot \mathbf{1}_{X_2 = +}(\mathbf{x}_2) \ge 0 \Big\}$$

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$$= \max\left\{\alpha \in \mathbb{R} : (1 - \varepsilon)E_p(\uparrow_{X_2} (f - \alpha \cdot \mathbf{1}_{X_1}) \cdot \mathbf{1}_{X_2 = +}) + \varepsilon \min_{\mathbf{x} \in \mathcal{X}} (f(\mathbf{x}_1) - \alpha) \cdot \mathbf{1}_{X_2 = +}(\mathbf{x}_2) \ge 0\right\}$$
$$= \max\left\{\alpha \ge \min_{x_1 \in \mathcal{X}_1} f(x_1) : (1 - \varepsilon) \sum_{x_1 \in \mathcal{X}_1} p_{(x_1, +)}(f(x_1) - \alpha) + \varepsilon \min_{x_1 \in \mathcal{X}_1} (f(x_1) - \alpha) \ge 0\right\}$$

$$\underline{E}^{X_1}(f|X_2=+) = \max\Big\{\alpha \in \mathbb{R} : \underline{E}_{p,\varepsilon}(\uparrow_{X_2} (f - \alpha \cdot \mathbf{1}_{X_1}) \cdot \mathbf{1}_{X_2=+}) \ge 0\Big\}$$

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$$= \max\left\{\alpha \geq \min_{x_1 \in \mathcal{X}_1} f(x_1) : (1 - \varepsilon) \frac{5}{9} \sum_{x_1 \in \mathcal{X}_1} p_{x_1}^{X_1|+} (f(x_1) - \alpha) + \varepsilon (\min_{x_1 \in \mathcal{X}_1} f(x_1) - \alpha) \geq 0 \right\}$$

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$$= \max \Big\{ \alpha \geq \min_{x_1 \in \mathcal{X}_1} f(x_1) : \alpha \leq (1-\delta) E_{p^{X_1|+}}(f) + \delta \min_{x_1 \in \mathcal{X}_1} f(x_1) \Big\}$$

with
$$\delta = rac{arepsilon}{(1-arepsilon)rac{5}{9}+arepsilon}$$

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$$= \underline{E}_{p^{X_1|+}, \delta}(f) \qquad \qquad \text{with } \delta = \frac{\varepsilon}{(1 - \varepsilon)\frac{5}{9} + \varepsilon}$$

Epistemic irrelevance & Epistemic independence

Consider $\boldsymbol{X} = (X_1, X_2)$ and a joint lower expectation $\underline{E}^{\boldsymbol{X}}$ on $(\mathcal{X} \to \mathbb{R})$

Epistemic irrelevance

- ▶ X_2 is epistemically irrelevant to X_1 iff $\underline{E}^{X_1}(\cdot|X_2 = x_2) := \underline{E}^{X_1}$ for all $x_2 \in \mathcal{X}_2$
- Epistemic irrelevance of X_2 to X_1 does not imply epistemic irrelevance of X_1 to X_2

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Epistemic independence

 X_1 and X_2 are called *epistemically independent* iff

- X_2 is epistemically irrelevant to X_1
- X_1 is epistemically irrelevant to X_2
Epistemically independent natural extension

Setup

- Consider two random variables X_1 and X_2
- ▶ Marginal lower expectations \underline{E}^{X_1} on $\mathcal{F}_1 \subseteq \mathcal{L}_1$ and \underline{E}^{X_2} on $\mathcal{F}_2 \subseteq \mathcal{L}_2$

Natural extension of epistemically independent marginals

Creating a *joint* lower expectation from epistemically independent marginals is not done as typically with a product

Epistemically independent natural extension

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Natural extension of epistemically independent marginals

- Creating a *joint* lower expectation from epistemically independent marginals is not done as typically with a product
- ▶ Natural extension of the implied assessment is used:

$$\underline{E}^{\mathbf{X}}(f) = (\underline{E}^{X_1} \boxtimes \underline{E}^{X_2})(f) := \sup_{g_1, g_2 \in \mathcal{L}} \min_{\mathbf{x} \in \mathcal{X}} \left(f(\mathbf{x}) - \left(g_1(\mathbf{x}) - \underline{E}_*^{X_1}(g_1(X_1, x_2)) \right) - \left(g_2(\mathbf{x}) - \underline{E}_*^{X_2}(g_2(x_1, X_2)) \right) \right)$$
for all $f \in \mathcal{L}$

Introduction to imprecise probabilities SIPTA School 2024, Ghent, Belgium

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12 August 2024